

APPENDIX A: ADDITIONAL TABLES

Table A1. List of industries and the corresponding industry classifications

JSIC code	Industries based on Japan Standard Industrial Classification (JSIC)	Industries based on the classification of the input–output table of the Chubu region
E	Construction	Construction
F-12	Manufacture of food	} Food
F-13	Manufacture of beverages, tobacco, and feed	
F-14	Manufacture of textile mill products	} Textile
F-15	Manufacture of apparel and other finished products	
F-16	Manufacture of lumber and wood products	Lumber
F-17	Manufacture of furniture and fixtures	Furniture
F-18	Manufacture of pulp, paper, and paper products	Pulp and paper
F-19	Printing and allied industries	Printing
F-20	Manufacture of chemical and allied products	Chemical
F-21	Manufacture of petroleum and coal products	Petroleum and coal
F-22	Manufacture of plastic products	Plastic
F-23	Manufacture of rubber products	Rubber
F-24	Manufacture of leather tanning, leather products	Leather
F-25	Manufacture of ceramic, stone, and clay products	Ceramic
F-26	Manufacture of iron and steel	Iron and steel
F-27	Manufacture of non-ferrous metals and products	Non-ferrous metals
F-28	Manufacture of fabricated metal	Fabricated metal
F-29	Manufacture of general machinery	General machinery
F-30	Manufacture of electrical machinery	Electrical
F-31	Manufacture of transportation equipment	Transportation
F-32	Manufacture of precision instruments and	Precision
G	Electricity, gas, heat supply, and water	Utilities
H	Transport	Transport
I	Wholesale and retail trade	Wholesale and retail
J	Finance and insurance	Finance and insurance
K	Real estate	Real estate
L	Services	} Information and communications Education and research Medical, health care and welfare Business services Personal services Public services

Note: The industry codes are based on the JSIC as of 1994.

Table A2. Significant local Moran statistics for the centre of growth clusters

Regional Industry	Local spatial statistics and associations based on the extensive spatial weights		Local spatial statistics and associations based on the geographical spatial weights	
The centre of the significant HH growth cluster:				
Toyota-shi in West Mikawa				
Food	n.s.		0.024	HH
Textile	n.s.		0.156	HH
Plastic	0.300	HH	0.126	HH
Rubber	0.130	HH	0.021	HH
Ceramic	0.015	HH	0.010	HH
Fabricated Metals	0.218	HH	0.191	HH
General Machinery	0.287	HH	0.198	HH
Electrical Machinery	-0.035	LH	-0.017	LH
Transportation	1.739	HH	0.406	HH
Construction	0.034	HH	0.023	HH
Wholesale and retail	0.358	HH	0.163	HH
Finance and insurance	0.059	HH	0.076	HH
Real Estate	n.s.		0.009	HH
Transport	0.211	HH	0.213	HH
Education and research	-0.023	LH	-0.006	LH
Medical	-0.186	LH	-0.050	LH
Public Services	0.021	HH	0.021	HH
Business Services	0.323	HH	0.484	HH
Personal Services	n.s.		-0.023	LH
Okazaki-shi in West Mikawa				
Food	n.s.		0.076	HH
Textile	n.s.		-0.040	LH
Furniture	n.s.		0.023	HH
Chemical	0.050	HH	0.037	HH
Plastic	0.027	HH	0.011	HH
Rubber	-0.019	LH	-0.003	LH
Ceramic	0.033	HH	0.020	HH
Fabricated Metals	0.092	HH	0.070	HH
General Machinery	0.182	HH	0.142	HH
Electrical Machinery	0.097	HH	0.054	HH
Transportation	1.352	HH	0.260	HH
Construction	n.s.		0.118	HH
Utilities	0.008	HH	0.007	HH
Wholesale and retail	-0.188	LH	-0.113	LH
Finance and insurance			0.023	HH
Real Estate	n.s.		0.026	HH
Transport	0.091	HH	0.095	HH
Education and research	0.095	HH	0.022	HH
Medical	n.s.		0.017	HH
Public Services			0.024	HH
Business Services	0.110	HH	0.129	HH
Personal Services	n.s.		-0.010	LH

Note: n.s. – Not significant. The significant values even at the 5% Bonferroni bound are shown in bold.

APPENDIX B: ESTIMATES OF OBSERVATIONS FOR ESDA

The error-component model employed is:

$$y(r, i, t) = comp(r) + comp(i) + comp(r, i) \\ + comp(t) + comp(i, t) + comp(r, t) + \varepsilon(r, i, t)$$

where:

$y(r, i, t)$: the average annual growth rate of employment in industry i in region r at time period t ,

$comp(r)$: time invariant regional trend component that is shared by industry,

$comp(i)$: time invariant sectoral trend component that is shared by all regions,

$comp(r, i)$: time invariant effect that is specific to industry i in region r ,

$comp(t)$: a pure time effect,

$comp(i, t)$: the interaction between a fixed industry and time effect,

$comp(r, t)$: the interaction between a fixed region and time effect,

$\varepsilon(r, i, t)$: an idiosyncratic disturbance that is orthogonal to all other effects.

The given dummy structure associated with each component is composed of terms with perfect collinearity, and is unidentified unless a sufficient number of restrictions are imposed. Here the following restrictions are imposed, taking the respective sample means as a reference point (GREEN and SEAKS, 1991; MARIMON and ZILLIBOTTI, 1998):

$$\left\{ \begin{array}{l} \sum_r comp(r) = 0, \\ \sum_t comp(t) = 0, \\ \sum_r comp(r, i) = 0, \forall i = 1, \dots, N_i, \\ \sum_i comp(r, i) = 0, \forall r = 1, \dots, N_r, \\ \sum_i comp(i, t) = 0, \forall t = 1, \dots, N_t, \\ \sum_t comp(i, t) = 0, \forall i = 1, \dots, N_i, \\ \sum_r comp(r, t) = 0, \forall t = 1, \dots, N_t, \\ \sum_t comp(r, t) = 0, \forall r = 1, \dots, N_r. \end{array} \right.$$

A set of $2N_r + 2N_i + 2N_t + 2$ restrictions, of which all but three restrictions are independent, guarantees the precise identification. Furthermore, imposing the above restrictions makes it possible to disentangle the dependence of $y(r, i, t)$ on the orthogonal components of each effect.

The absolute changes in growth rates of the small regional sectors tend to be considerably larger than those in the large sectors, and this results in an inherent heteroscedasticity problem in the ordinary least squares estimation. To address this issue, the estimation model is weighted by a factor given by the employment of each regional industry divided by aggregate employment in all industries at time period t . This weight also reflects their respective importance for aggregate employment (SUEDEKUM and BLIEN, 2005).

The (weighted) long-term growth across regional industries $y(r, i)$ is captured by the

following components, which are not affected by the industry-mix effect or any time-specific effects:

$$y(r, i) = comp(r) + comp(r, i)$$

The estimated values of $y(r, i)$ are used as the observations for ESDA.

APPENDIX C: GLOBAL AND LOCAL STATISTICS ON SPATIAL ASSOCIATIONS

The global spatial association for industry i 's adjusted long-term growth in region r , y_{ri} , can be formally captured by using the following extended version of Moran's I

$$I(d, f) = \frac{\sum_r \sum_i \sum_s \sum_j [w_{ri,sj}^*(d, f)] (y_{ri} - \bar{y})(y_{sj} - \bar{y})}{\sum_r \sum_i (y_{ri} - \bar{y})^2}$$

where $w_{ri,sj}^*(d, f)$ denotes the element of the row-standardized extensive spatial weight matrix. Consequently, the expected Moran's I value can be given by $E(I) = -1/(N_r N_i - 1)$.

The local spatial associations of the adjusted long-term industrial growth are measured with the following extended version of the local Moran statistic:

$$I_{ri}(d, f) = \frac{y_{ri} - \bar{y}}{m_1} \sum_s \sum_j [w_{ri,sj}^*(d, f)] (y_{sj} - \bar{y})$$

where:

$$m_1 = \sum_r \sum_i \frac{(y_{ri} - \bar{y})^2}{N_r N_i}$$

REFERENCES

- GREEN W. H. and SEAKS T. G. (1991) The restricted least squares estimator: A pedagogical note, *Review of Economics and Statistics* 73, 563–567.
- MARIMON R. and ZILIBOTTI F. (1998) 'Actual' and 'virtual' employment in Europe: Is Spain different?, *European Economic Review* 42, 123–153.
- SUEDEKUM J. and BLIEN U. (2005) Local economic structure and industry development in Germany, 1993-2001, *Economics Bulletin* 15, 1–8.